Estimation of a fold convolution in additive noise model with compactly supported noise density

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Abstract – Consider the model \( Y = X + Z \), where \( Y \) is an observable random variable, \( X \) is an unobservable random variable with unknown density \( f \), and \( Z \) is a random noise independent of \( X \). The density \( g \) of \( Z \) is known exactly and assumed to be compactly supported. We are interested in estimating the \( m \)-fold convolution \( f_m = f \ast \cdots \ast f \) on the basis of independent and identically distributed (i.i.d.) observations \( Y, \ldots, Y_n \) drawn from the distribution of \( Y \). Based on the observations as well as the ridge-parameter regularization method, we propose an estimator for the function \( f_n \) depending on two regularization parameters in which a parameter is given and a parameter must be chosen. The proposed estimator is shown to be consistent with respect to the mean integrated squared error under some conditions of the parameters. After that we derive a convergence rate of the estimator under some additional regular assumptions for the density \( f \).

Index Terms – estimator, compactly supported noise density, convergence rate

1 INTRODUCTION

In this paper, we consider the additive noise model

\[ Y = X + Z \quad \text{(1)} \]

where \( Y \) is an observable random variable, \( X \) is an unobservable random variable with unknown density \( f \), and \( Z \) is an unobservable random noise with known density \( g \). The density \( g \) is called noise density. We also suppose that \( X \) and \( Z \) are independent. Estimating \( f \) on basis of i.i.d. observations of \( Y \) has been known as the density deconvolution problem in statistics. This problem has received much attention during two last decades. Various estimation techniques for \( f \) can be found in Carroll-Hall [1], Stefanski-Carroll [2], Fan [3], Neumann [4], Pensky-Vidakovic [5], Hall-Meister [6], Butucea-Tsybakov [7], Johannes [8], among others.

This problem has concerned with many real-life problems in econometrics, biometrics, signal reconstruction, etc. For example, when an input signal passes through a filter, output signal is usually disturbed by an additional noise, in which the output signal is observable, but the input signal is not.

Let \( Y, \ldots, Y_n \) be \( n \) i.i.d. observations of \( Y \).

In the present paper, instead of estimating \( f \), we focus on the problem of estimating the \( m \)-fold convolution

\[ f_m = f \ast \cdots \ast f, \quad m \in \mathbb{Z}^+, \quad \text{m times} \quad \text{(2)} \]

based on the observations. In the free-error case, i.e. \( Z = 0 \), there are many papers related to this problem, such as Frees [9], Saavedra-Cao [10], Ahmad-Fan [11], Ahmad-Mugdadi [12], Chesneau et al. [13], Chesneau-Navarro [14], and references therein. For \( m = 1 \), the problem of estimating \( f \) reduces to the density deconvolution problem. To the best of our knowledge, for \( m \in \mathbb{Z}^+, m \geq 2 \), so far this problem has been only studied by Chesneau et al. [15]. In that paper, the authors constructed a kernel type of estimator for \( f_n \) under the assumption that \( g^n \) is nonvanishing on \( \mathbb{R} \), where the function \( g^n(t) = \int_{\mathbb{R}} f(x) e^{-int} \, dx \) is the Fourier transform of \( g \). The latter assumption is fulfilled with many usual densities, such as normal, Cauchy, Laplace, gamma, chi-square densities. However, there are also several cases of \( g \) that cannot be applied to this paper. For

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instance, the case in which \( g \) is a uniform density or a compactly supported density in general. In the present paper, as a continuation of the paper of Chesneau et al. [15], we consider the case of compactly supported noise density \( g \). In fact, the problem was studied by Trong-Phuong [16] in the case of \( m = 1 \); however, the problem has more challenge with \( m \in \mathbb{Z}^+ \), \( m \geq 2 \).

The rest of our paper consists of three sections. In Section 2, we establish our estimator. In Section 3, we state main results of our paper. Finally, some conclusions are presented in Section 4.

For convenience, we introduce some notations. For two sequences \( (u_n) \) and \( (v_n) \) of positive real numbers, we write \( u_n = O(v_n) \) if the sequence \( (u_n / v_n) \) is bounded. The number of \( k \)-combinations from a set of \( p \) elements is denoted by \( \binom{p}{k} \). The number \( \lambda(A) \) is the Lebesgue measure of a set \( A \subseteq \mathbb{R} \). For a function \( \varphi \in L^p(\mathbb{R}), \ 1 \leq p \leq \infty \), the symbol \( \|\varphi\|_p \) represents the usual \( L^p(\mathbb{R}) \)-norm of \( \varphi \). For a function \( \phi : \mathbb{R} \to \mathbb{C} \), we define \( Z(\phi) = \{x \in \mathbb{R} : \phi(x) = 0\} \) and \( \text{supp}(\phi) = \mathbb{R} \setminus Z(\phi) \), the closure in \( \mathbb{R} \) of the set \( \mathbb{R} \setminus Z(\phi) \). Regarding the Fourier transform, we recall that \( \left( \phi^n \right)(x) = 2\pi \phi(-x) \) for \( x \in \mathbb{R} \), \( \phi \in L^1(\mathbb{R}) \cap L^2(\mathbb{R}) \), and moreover, \( \|\phi^n\|_2 = \sqrt{2\pi} \|\phi\|_1 \), which is called the Parseval identity.

2 METHODS

We now describe the method for constructing an estimator for \( f_m \). First, from the equation (2) we have \( f_m^n(t) = [f^n(t)]^m \). Also, from the independence of \( X \) and \( Z \), we obtain \( h = f * g \), where \( h \) is density of \( Y \). The latter equation gives \( h^n(t) = f^n(t) \cdot g^n(t) \), so \( f_m^n(t) = [h^n(t) / g^n(t)]^m \) if \( g^n(t) \neq 0 \). Then applying the Fourier inversion formula, we can obtain an estimator for \( f_m \). However, it is very dangerous to use \([h^n(t) / g^n(t)]^m \) as an estimator for \( f_m^n(t) \) in case \( g^n \) can vanish on \( \mathbb{R} \). In this case, to avoid division by numbers very close to zero, \( 1/ g^n(t) \) is replaced by the quantity \( r(t) = g^n(-t) / \max \{g^n(t)^2 ; \delta|t| \} \), called the ridge function. Here \( a > 1/m \) is a given parameter, and \( \delta > 0 \) is a regularization parameter that will be chosen according to \( n \) later so that \( \delta \to 0 \) as \( n \to \infty \). We then obtain an estimator for \( f_m^n(t) \) in the form \( \hat{\Phi}(t) = [r(t)h^n(t)]^m \).

Nevertheless, the function \( \Phi(t) \) depends on the Fourier transform \( h^n(t) \), which is an unknown quantity, and so, we cannot use \( \Phi(t) \) to estimate \( f_m^n(t) \). Fortunately, from the i.i.d. observations \( Y_1, \ldots, Y_n \), we can estimate \( h^n(t) \) by the empirical characteristics function \( \hat{h}^n(t) = n\sum_{i=1}^n e^{itY_i} \).

Hence, another estimator for \( f_m^n(t) \) is proposed by \( \Phi(t) = \left[ r(t)\hat{h}^n(t) \right]^m \). Finally, using the Fourier inversion formula, we derive an estimator for \( f_m \) in the final form

\[
\hat{f}_{m, \delta}(x) = \frac{1}{2\pi} \int_{\mathbb{R}} e^{-itx} \Phi(t) dt
= \frac{1}{2\pi} \int_{\mathbb{R}} e^{-itx} \left( \frac{g^n(-t)\hat{h}^n(t)}{\max \{g^n(t)^2 ; \delta|t| \}} \right)^m dt. \tag{3}
\]

Note that the condition \( a > 1/m \) implies \( \Phi \in L^1(\mathbb{R}) \cap L^2(\mathbb{R}) \) almost surely. Thus, the estimator \( \hat{f}_{m, \delta}(x) \) is well-defined for all values of \( x \in \mathbb{R} \), and moreover, \( \hat{f}_{m, \delta} \) belongs to \( L^1(\mathbb{R}) \).

3 RESULTS

In this section, we consider consistency and convergence rate of the estimator \( \hat{f}_{m, \delta} \) given in (3) under the mean integrated squared error \( \text{MISE}(\hat{f}_{m, \delta}, f_m) = \mathbb{E} \left\| \hat{f}_{m, \delta} - f_m \right\|^2 \). First, a general bound for \( \text{MISE}(\hat{f}_{m, \delta}, f_m) \) is given in the following proposition.

**Proposition 1.** Let \( \hat{f}_{m, \delta}, \ m \geq 1, \ be \ as \ in \ (3) \ with \ a > 1/m \ and \ 0 < \delta < 1. \ Suppose \ that \ f \in L^2(\mathbb{R}). \ Then we have**
MISE(\hat{f}_{m, \delta}, f_\infty) \leq \frac{1}{\pi} \int_\mathbb{R} \left| \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right|^\beta \left| f^n(t) \right|^\alpha dt \\
+ \frac{2^{n-1}}{\pi} \sum_{k=1}^{\infty} C_k \int_\mathbb{R} \left| \frac{|g^{n-k}(t)|^{\alpha-k} |f^n(t)|^{\alpha-2k}}{\max \{|g^{n-k}(t)|^{\alpha-k}; |\delta|^\gamma\}} \right|^\beta \left| (f^n(t))^{\alpha-k} \right|^{\alpha-k} dt,

where $C_k = (72k)^{\frac{2}{n-k}} (2k/(2k-1))^k$, $k = 1, \ldots, m$.

**Proof.** Since $f$ is a density and is in $L^2(\mathbb{R})$, we deduce $f_\infty \in L^1(\mathbb{R}) \cap L^2(\mathbb{R})$, so $f_\infty \in L^\infty(\mathbb{R})$.

Using the Parseval identity, the Fubini theorem and the binomial theorem, we obtain

$$
\begin{align*}
\text{MISE}(\hat{f}_{m, \delta}, f_\infty) &= \frac{1}{\pi} \int_\mathbb{R} \left| \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right|^\beta \left| f^n(t) \right|^\alpha dt \\
&= \frac{1}{\pi} \int_\mathbb{R} \left| \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right|^\beta \left| f^n(t) \right|^\alpha dt \\
&= \frac{1}{\pi} \int_\mathbb{R} \left[ \left( \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right)^\beta - \left| f^n(t) \right|^\alpha \right] dt \\
&= \frac{1}{\pi} \int_\mathbb{R} \left[ \left( \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right)^\beta \sum_{k=1}^{\infty} C_k \left( \left( \frac{|g^{n-k}(t)|^{\alpha-k}}{\max \{|g^{n-k}(t)|^{\alpha-k}; |\delta|^\gamma\}} \right)^\beta \right) \left| f^n(t) \right|^{\alpha-k} - \left| f^n(t) \right|^\alpha \right] dt.
\end{align*}
$$

Using the inequality $|z_1 + z_2|^\alpha \leq 2 |z_1|^\alpha + 2 |z_2|^\alpha$ with $(z_1, z_2) \in \mathbb{C}^2$ yields

$$
\text{MISE}(\hat{f}_{m, \delta}, f_\infty) \leq \frac{1}{\pi} \left[ I_1 + I_2 \right],
$$

where

$$
I_1 = \int_\mathbb{R} \left| \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right|^\beta \left| f^n(t) \right|^\alpha \left| f^n(t) \right|^\alpha dt,
$$

$$
I_2 = \int_\mathbb{R} \left[ \left( \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right)^\beta - \left| f^n(t) \right|^\alpha \right] \sum_{k=1}^{\infty} C_k \left( \left( \frac{|g^{n-k}(t)|^{\alpha-k}}{\max \{|g^{n-k}(t)|^{\alpha-k}; |\delta|^\gamma\}} \right)^\beta \right) \left| f^n(t) \right|^{\alpha-k} dt.
$$

Since $h^n(t) = f^n(t) \cdot g^n(t)$ and $g^n(-t) = \overline{g^n(t)}$, in which $\overline{g^n(t)}$ denotes the conjugate of $g^n(t)$, we have

$$
\begin{align*}
&I_1 = \int_\mathbb{R} \left| \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right|^\beta \left| f^n(t) \right|^\alpha \left| f^n(t) \right|^\alpha dt \\
&= \int_\mathbb{R} \left[ \left( \frac{|g^n(t)|^\alpha}{\max \{|g^n(t)|^\alpha; |\delta|^\gamma\}} \right)^\beta \sum_{k=1}^{\infty} C_k \left( \left( \frac{|g^{n-k}(t)|^{\alpha-k}}{\max \{|g^{n-k}(t)|^{\alpha-k}; |\delta|^\gamma\}} \right)^\beta \right) \left| f^n(t) \right|^{\alpha-k} \right] dt.
\end{align*}
$$

Define $U_j = n^{-1} \left[ e^{\alpha t_j} - \mathbb{E}(e^{\alpha t_j}) \right]$, $j = 1, \ldots, n$.

Clearly, the sequence $(U_j)_{j=1, \ldots, n}$ satisfies the conditions of Lemma A.1 in Chesneau et al. [15], and moreover, $|U_j| \leq 2/n$. Hence, applying Lemma A.1 in Chesneau et al. [15] with $p = 2k > 1$, we get

$$
\mathbb{E} \left[ \left\| \sum_{j=1}^{n} U_j \right\|^\beta \right] \leq (36k)^{\frac{2k}{2k-1}} \left( \frac{2k}{2k-1} \right)^{\frac{4}{n}} \left( \frac{4}{n} \right) = (72k)^{\frac{2k}{2k-1}} \left( \frac{2k}{2k-1} \right) \cdot \frac{1}{n} = C_n \frac{1}{n}.
$$

Thus,

$$
I_2 \leq (2^{n-1} \sum_{k=1}^{\infty} C_k \int_\mathbb{R} \left| \frac{|g^{n-k}(t)|^{\alpha-k}}{\max \{|g^{n-k}(t)|^{\alpha-k}; |\delta|^\gamma\}} \right|^\beta \left| f^n(t) \right|^{\alpha-k} dt. \tag{6}
$$

From (4) – (6), we obtain the desired conclusion.

**Proposition 2.** Let the assumptions of Proposition 1 hold. Then there exists a $k_0 > 0$ depending only on $g$ such that
\[
\sum_{k=1}^{n} C_{k} \int_{\mathbb{R}} \frac{|g^{n}(t)|^{1/n-2k} |f^{\alpha}(t)|^{2(n-k)}}{\max\left\{ |g^{n}(t)|^{2}; |\delta f(t)|^{2} \right\} \ dt}
\]
\[
\leq 2^{2n+1} k_{0} + \int_{|\delta f|} \frac{1}{|\delta f|} \ dt \left( \sum_{k=1}^{n} C_{k} \right) \frac{1}{n \delta^{n}}.
\]

**Proof.** Since the function \( |g^{n}| \) is continuous on \( \mathbb{R} \) and \( |g^{n}(0)| = 1 \), there is a constant \( k_{0} > 0 \) depending only on \( g \) such that \( |g^{n}(t)| > 1/2 \) for all \( |t| \leq k_{0} \). Then for \( k = 1, \ldots, m \) we have
\[
\int_{\mathbb{R}} \frac{|g^{n}(t)|^{1/n-2k} |f^{\alpha}(t)|^{2(n-k)}}{\max\left\{ |g^{n}(t)|^{2}; |\delta f(t)|^{2} \right\} \ dt
\]
\[
= \int_{\mathbb{R}} \frac{|g^{n}(t)|^{2n}}{\max\left\{ |g^{n}(t)|^{2}; |\delta f(t)|^{2} \right\}^{2n} \ dt
\]
\[
\leq \int_{|\delta f|} \frac{1}{|\delta f|^{2n}} \ dt + \int_{|\delta f|} \frac{1}{\max\left\{ |g^{n}(t)|^{2}; |\delta f(t)|^{2} \right\}^{2n}} \ dt
\]
\[
\leq 2^{2n+1} k_{0} + \int_{|\delta f|} \frac{1}{|\delta f|} \ dt \left( \sum_{k=1}^{n} C_{k} \right) \frac{1}{n \delta^{n}}.
\]
Hence,
\[
\sum_{k=1}^{n} C_{k} \int_{\mathbb{R}} \frac{|g^{n}(t)|^{1/n-2k} |f^{\alpha}(t)|^{2(n-k)}}{\max\left\{ |g^{n}(t)|^{2}; |\delta f(t)|^{2} \right\}^{2n} \ dt
\]
\[
\leq 2^{2n+1} k_{0} + \int_{|\delta f|} \frac{1}{|\delta f|} \ dt \left( \sum_{k=1}^{n} C_{k} \right) \frac{1}{n \delta^{n}}.
\]
The proof of the proposition is completed.

The mean consistency of the estimator \( \hat{f}_{m,\delta} \) is given in the following theorem.

**Theorem 3.** Suppose that \( f \in L^{1}(\mathbb{R}) \) and \( \lambda(\mathcal{Z}(g^{n})) = 0 \). Let \( \hat{f}_{m,\delta} \) be as in (3), where \( \alpha > 1/m \) and \( \delta \) is a positive parameter depending on \( n \) such that \( \delta \to 0 \) and \( n \delta^{m} \to \infty \) as \( n \to \infty \). Then \( \text{MISE}(\hat{f}_{m,\delta}, f) \to 0 \) as \( n \to \infty \).
\[ \mathcal{F}(M) = \{ \text{density } \phi \text{ on } \mathbb{R} : \phi \in L^2(\mathbb{R}), \supp(\phi) \subseteq [-M, M] \}, \]

\[ \mathcal{F}(c_1, c_2, d, \beta) = \{ \text{density } \phi \text{ on } \mathbb{R} : c_1 e^{-\alpha \mathcal{L}} \leq |\phi^\prime(t)| \leq c_2 e^{-\alpha \mathcal{L}}, \forall t \in \mathbb{R} \}, \]

in which \( M, c_1, c_2, d, \beta \) are positive constants.

The class \( \mathcal{F}(M) \) includes compactly supported densities on \([-M, M]\). The class \( \mathcal{F}(c_1, c_2, d, \beta) \) contains densities in which Fourier transforms converge to zero with exponential rate of order \( \beta \). Normal and Cauchy densities are typical examples of \( \mathcal{F}(c_1, c_2, d, \beta) \). In fact, using the Fourier inversion formula and the Lebesgue dominated convergence theorem, one can show that each element of \( \mathcal{F}(c_1, c_2, d, \beta) \) is an infinitely differentiable function on \( \mathbb{R} \). Hence, \( \mathcal{F}(c_1, c_2, d, \beta) \) is often called the class of “super-smooth” densities. In fact, the case of \( g \in \mathcal{F}(c_1, c_2, d, \beta) \) has been studied in Chesneau et al. [15]. The reason for considering this class in the present paper is that we want to demonstrate that the estimator \( \hat{g}_{n, \delta} \) can also be attained the convergence rate established in Chesneau et al. [15].

Now, we consider the case \( g \in \mathcal{F}(M) \). Before stating main result of our paper, we need the following auxiliary lemma. This auxiliary lemma is not a new result. It is quite similar to Theorem 3 in Trong-Phuong [16].

**Lemma 5.** Suppose \( g \in \mathcal{F}(M) \). Given \( \mu > 1 \). For \( \rho > 0 \) small enough, we choose an \( R > 0 \) depending on \( \rho \) such that

\[ 2eMR[(\mu + 1)\ln R + \ln(15e^\rho)] = \ln(\rho^{-1}). \]

Then for \( \rho > 0 \) small enough we have

\[ 30(\mu + 1)Me^\rho \leq R \leq (2eM)^{-1} \ln(\rho^{-1}) \]

In addition, we have \( \lambda(B_{R, \rho}) \leq 2R^{-\mu} \), where

\[ B_{R, \rho} = \{ t \in \mathbb{R} : |t| \leq R, |g^\delta(t)| \leq \rho \}. \]

Main result of our paper is the following theorem.

**Theorem 6.** Let \( \alpha > 1/2, \ L > 0 \). Assume that \( g \in \mathcal{F}(M) \) with \( M > 0 \). Let \( \hat{f}_{n, \delta} \) be as in (3) for a known \( a > 1/m \) and \( \delta = n^{-\tau} \) with \( 0 < \tau < 1/m \). Then we have

\[ \sup_{f \in \mathcal{F}(\alpha, L)} \text{MISE}(\hat{f}_{n, \delta}, \hat{f}_n) = O\left( (\ln n)^{-m - 1} \right). \]

**Proof.** Suppose \( f \in \mathcal{F}(\alpha, L) \). We take \( \mu = 2m\alpha > 1, \ 0 < \tau < 1/m \) and \( \rho = n^{-\tau} \) with \( 0 < \nu < \tau / 2 \). Then applying Lemma 5 gives that there exists an \( R > 0 \) depending on \( n \) such that

\[ \left[ 30(2m\alpha + 1)Me^\rho \right]^{1/2} \left[ \ln(n^\nu) \right]^{1/2} \leq R \leq (2eM)^{-1} \ln(n^\nu), \]

and \( \lambda(B_{R, \rho}) \leq 2R^{-2m\alpha} \) for \( n \) large enough. Now, for \( 0 < \delta < R^{-\mu} \), we have

\[ J := \int_{\mathbb{R}} \frac{g^\delta(t)}{\max\{g^\delta(t), |\delta|^2\}} - 1 \left| f^\delta(t) \right|^2 dt \]

\[ \leq \int_{|t| \leq R} \left| f^\delta(t) \right|^2 dt + \int_{|t| > R} \left| f^\delta(t) \right|^2 dt \]

\[ \leq \lambda(B_{R, \rho}) + \int_{|t| > R} \left| f^\delta(t) \right|^2 dt \leq 2R^{-2m\alpha} + \int_{|t| > R} \left| f^\delta(t) \right|^2 dt, \]

where we note that \( \{ t \in \mathbb{R} : |t| \leq R, \rho < |g^\delta(t)| < \sqrt{|\delta|^2} \} = \emptyset \). Moreover, since \( f \in \mathcal{F}(\alpha, L) \), we derive

\[ \int_{|t| \leq R} \left| f^\delta(t) \right|^2 dt \]

\[ \leq (2 + L^\nu) \int_{|t| > R} \left[ f^\delta(t)^2 + (1 + t^2)^\nu \right] \left[ f^\delta(t)^2 + (1 + t^2)^\nu \right] dt \]

\[ \leq L^\nu R^{-2m\alpha} \]

Hence,

\[ J \leq (2 + L^\nu) R^{-2m\alpha} \]

\[ \leq (2 + L^\nu) \left[ 30(2m\alpha + 1)Me^\rho \right]^{1/2} \left[ \ln(n^\nu) \right]^{-m - 1} \]

\[ = O\left( (\ln n)^{-m - 1} \right). \]

Combining (7) with Proposition 1 and Proposition 2, we obtain

\[ \text{MISE}(\hat{f}_{n, \delta}, f_n) = O\left( (\ln n)^{-m - 1} + (n\delta^{m - 1})^{-1} \right). \]

Now, we need to choose \( \delta > 0 \) according to \( n \) so that \( \delta < R^{-\mu} \rho \), and rate of convergence of \( (n\delta^{m - 1})^{-1} \) is faster than that of \( (\ln n)^{-m - 1} \). A possible choice is...
Suppose that $\delta = n^{-r}$. Then the conclusion of the theorem is followed.

**Remark 7.** The parameter $\delta$ in Theorem 6 does not depend on $\alpha$, the prior degree of smoothness of $f$. Therefore, the estimator $\hat{f}_{m,\delta}(x)$ can be computed with out any knowledge concerning the degree of smoothness.

Finally, we consider the case $g \in \mathcal{F}(c_1, c_2, d, \beta)$. We have

**Theorem 8.** Let $\alpha > 1/2$, $L > 0$. Suppose that $g \in \mathcal{F}(c_1, c_2, d, \beta)$, where $c_1, c_2, d, \beta$ are the given positive constants. Let $\hat{f}_{m,\delta}$ be as in (3) for a known $\alpha > 1/m$ and $\delta = n^{(1-8m)(1+4m)}(\ln n)^{(1-8m)/(4m\beta)}$. Then we have

$$\sup_{f \in \mathcal{F}(\alpha, L)} \text{MISE}(\hat{f}_{m,\delta}, f_m) = O\left((\ln n)^{-2m/\alpha}/\alpha\right).$$

**Proof.** Suppose $f \in \mathcal{F}(\alpha, L)$. Let $T$ be a positive number that will be selected later. Using the inequality

$$0 \leq \max\left\{|g^n(t)|^2; \delta|f'|^2\right\}^m - |g^n(t)|^{2m} \leq \delta^m |f'|^{2m}$$

for all $t \in \mathbb{R}$ and the assumptions $g \in \mathcal{F}(c_1, c_2, d, \beta), f \in \mathcal{F}(\alpha, L)$, we have

$$Q := \sum_{k=1}^{n} C_k^i \left[ \int_{\mathbb{R}} \left| \frac{g^n(t)}{\max\left\{|g^n(t)|^2; \delta|f'|^2\right\}} \right|^m f^n(t) \left| \frac{|f^n(t)|^2}{\delta^m} \right|^{2m} dt \right].$$

$$= \sum_{k=1}^{n} C_k^i \int_{\mathbb{R}} \left| \frac{g^n(t)}{\max\left\{|g^n(t)|^2; \delta|f'|^2\right\}} \right|^m f^n(t) \left| \frac{|f^n(t)|^2}{\delta^m} \right|^{2m} dt \right].$$

For the quantities $Q_i$ and $Q_2$, we have the estimates

$$Q_1 \leq \sum_{k=1}^{n} C_k^i \left[ \frac{1}{n^5} \left| \frac{g^n(t)}{\max\left\{|g^n(t)|^2; \delta|f'|^2\right\}} \right|^m f^n(t) \left| \frac{|f^n(t)|^2}{\delta^m} \right|^{2m} dt \right].$$

$$Q_2 \leq \sum_{k=1}^{n} C_k^i \left[ \int_{\mathbb{R}} \left| \frac{f^n(t)}{g^n(t)} \right|^m f^n(t) \left| \frac{|f^n(t)|^2}{\delta^m} \right|^{2m} dt \right].$$

Combining Proposition 1 with the estimates of $J$, $Q_1$, and $Q_2$, we get

$$\text{MISE}(\hat{f}_{m,\delta}, f_m) = O\left(\frac{\delta^{2m} e^{4mT^2} + T^{2m}}{n^5 \delta^{2m}}\right).$$

Also,
Choosing \( T = \left( (\ln n) / (8m\delta) \right)^{1/\beta} \) yields

\[
\text{MISE} \left( \hat{f}_{m, \delta}, f_m \right) = O \left\{ n^{1/2} \left( \ln n \right)^{2\alpha + \beta} \delta^{2\alpha + \beta} + \left( \ln n \right)^{-2\alpha + \beta} \right\\ + \sum_{k=1}^{m-1} C_m \lambda \max \left\{ \left( c_1 \right)^{3-k} ; \left( c_2 \right)^{4-k} \right\} \times O \left\{ \left( n^{1-1/(4\alpha+1)} \right)^{2\alpha + \beta} + \left( \ln n \right)^{-2\alpha + \beta} \right\\ + \sum_{k=1}^{m-1} C_m \lambda \max \left\{ \left( c_1 \right)^{3-k} ; \left( c_2 \right)^{4-k} \right\} \right.
\]

Choosing \( \delta = n^{3-2\alpha-\beta}^{1/(4\alpha+1)} \left( \ln n \right)^{1/(4\alpha+1)} / \left( 4\alpha + \beta \right) \) implies the desired conclusion.

**Remark 9.** We see that the convergence rate of MISE \( \left( \hat{f}_{m, \delta}, f_m \right) \) uniformly over the class \( \mathcal{F}(\alpha, L) \) in Theorem 8 is as same as that of Chesneau et al. [15] when \( g \in \mathcal{F}(c_1, c_2, d, \beta) \). In particular, when \( m = 1 \), the convergence rate also coincides with the optimal rate of convergence proven in Fan [3].

4 CONCLUSIONS

We have considered the problem of nonparametric estimation of the \( m \)-fold convolution \( f_m \) in the additive noise model (1), where the noise density \( g \) is known and assumed to be compactly supported. An estimator for the function \( f_m \) has been proposed and proved to be consistent with respect to the mean integrated squared error. Under some regular conditions for the density \( f \) of \( X \), we derive a convergence rate of the estimator. We also have shown that the estimator attains the same rate as the one of Chesneau et al. [15] if the density \( g \) is supersmooth. A possible extension of this work is to study our estimation procedure in the case of unknown noise density \( g \). We leave this problem for our future research.

**REFERENCES**

Ước lượng một tự tích chập trong một mô hình công nhiều với hàm mật độ nhiễu có giá compact

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Tóm tắt – Bài báo này đề cập mô hình $Y = X + Z$, trong đó $Y$ là một biến ngẫu nhiên quan sát được, $X$ là một biến ngẫu nhiên không quan sát được với hàm mật độ $f$ chưa biết, và $Z$ là nhiễu ngẫu nhiên độc lập với $X$. Hàm mật độ $g$ của $Z$ được giả thiết biết chính xác và có giá compact. Bài báo nghiên cứu vấn đề ước lượng phi tham số cho tự tích chập $f_m = f * ... * f$ (m lần) trên cơ sở mẫu quan sát $Y_1, ..., Y_n$ độc lập, cùng phân phối được lấy từ phân phối của $Y$. Dựa trên các quan sát này cùng như phương pháp chính hóa tham số chọn, ước lượng được $f$ phù thuộc vào hai tham số chính hóa được đề xuất, trong đó một tham số được chọn trước và tham số còn lại sẽ được chọn sau. Ước lượng này được chứng tỏ là vùng tương ứng với trung bình sáng của tham số chính hóa của hàm mật độ $f$.

Từ khóa – Uớc lượng, hàm mật độ nhiễu có giá compact, tốc độ hội tụ